

## ASYMPTOTICALLY OPTIMAL ADAPTIVE CONTROL WITH CONSISTENT PARAMETER ESTIMATES\*

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**Abstract.** For the discrete-time linear stochastic systems with unknown coefficients we give an adaptive control by which both strong consistency of the parameter estimates and asymptotic optimality for the tracking system are achieved simultaneously. This is done by disturbing the signal that is to be tracked, and the disturbance consists of a sequence of random vectors with covariance matrices tending to zero. The main result is essentially based on some criteria for consistency of parameter estimate for the system without monitoring, which are also demonstrated in the paper. The existence of adaptive control is also discussed.

**Key words.** stochastic system, parameter estimate, strong consistency, adaptive tracking, asymptotic optimality

**AMS(MOS) subject classification.** 93C40

**1. Introduction.** Since Åström and Wittenmark [1] introduced the self-tuning regulator, much work has been devoted in recent years to the parameter-adaptive control and to the related parameter estimation problem. For the adaptive tracking problem, Goodwin, Ramadge and Caines [15] and Sin and Goodwin [21] have established the global convergence of the system and the asymptotic optimality of the tracking error by use of the stochastic gradient and the modified least squares algorithms, respectively. On the other hand, for linear stochastic systems without monitoring there are different conditions guaranteeing the strong consistency of estimates for the unknown system coefficients by invoking various approaches such as the probabilistic method (Ljung [18], Moore [20], Solo [22]), the ordinary differential equation method (Ljung [19], Kushner and Clark [17]) and the combined treatment (Chen [5], [6], [8]). But the crucial point in these different conditions is almost the same fact—the persistent excitation condition, which means that for the matrix  $\sum_{i=1}^n \varphi_i \varphi_i^T$  consisting of the stochastic regressors  $\varphi_i$  the ratio of its maximum to minimum eigenvalues is bounded. Unfortunately, it does not always take place for the system with asymptotically optimal adaptive control given in Goodwin, Ramadge and Caines [15] and Sin and Goodwin [21], as shown in Becker, Kumar and Wei [2].

In order to get the consistent estimate for unknown parameters the adaptive control law is disturbed by a random noise introduced artificially (see Caines and Lafortune [3], Chen [7], Chen and Caines [9]). With such a treatment it turns out that the estimate is strongly consistent but the tracking error differs from its minimal value by an additional term caused by the random noise added to the adaptive control law.

However, all these facts do not mean that there is no adaptive control law forcing the long run average of the tracking errors to be minimal and, at the same time, making the parameter estimate strongly consistent, since the asymptotically optimal adaptive control law is not unique.

In this paper, we first give an adaptive control by which both strong consistency of the estimates and optimality for the tracking system are achieved simultaneously. The main idea is that the asymptotically optimal adaptive control is disturbed by a random vector sequence with vanishing covariance matrices, in contrast to the work of Caines and Lafortune [3], Chen and Caines [9] and Chen [7], where the disturbance

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is of constant covariance matrix. As a result the matrix  $\sum_{i=1}^n \varphi_i \varphi_i^T$  mentioned above is ill-conditioned; hence no persistent excitation-like condition can be applied to guarantee consistency for estimates. However, recently the authors have obtained some new results (Chen and Guo [10], [11], [12]), establishing the strong consistency of parameter estimates for systems with  $\sum_{i=1}^n \varphi_i \varphi_i^T$  ill-conditioned, and it appears that they are suitable to the analysis of the case of adaptive control with vanishing disturbances and make the system asymptotically optimal and the parameter estimates strongly consistent.

**2. Statement of the problem.** Let  $(\Omega, \mathcal{F}, P)$  be a probability space with a family  $\{\mathcal{F}_n\}$  of nondecreasing sub- $\sigma$ -algebras. Consider the following stochastic control system:

$$(2.1) \quad y_n + A_1 y_{n-1} + \dots + A_p y_{n-p} = B_1 u_{n-1} + \dots + B_q u_{n-q} + w_n + C_1 w_{n-1} + \dots + C_r w_{n-r}$$

where  $y_n$ ,  $u_n$  and  $w_n$  are the  $m$ -,  $l$ - and  $m$ -dimensional output, input and driven noise, respectively, and  $p \geq 1$ ,  $q \geq 1$ ,  $y_n = w_n = 0$ ,  $u_n = 0$  for  $n < 0$ .  $A_i$ ,  $B_j$ ,  $C_k$  ( $i = 1 \dots p$ ,  $j = 1 \dots q$ ,  $k = 1 \dots r$ ) are the unknown matrices.

Assume that  $u_n$  and  $w_n$  are  $\mathcal{F}_n$ -measurable and

$$(2.2) \quad E(w_n | \mathcal{F}_{n-1}) = 0, \quad E(\|w_n\|^2 | \mathcal{F}_{n-1}) \leq c_0 r_{n-1}^\varepsilon$$

with constants  $c_0 > 0$ ,  $\varepsilon \in [0, 1)$  and  $r_{n-1}$  defined later on by (2.9).

Let  $z$  be the shift-back operator and set

$$(2.3) \quad A(z) = I + A_1 z + \dots + A_p z^p,$$

$$(2.4) \quad B(z) = B_1 + B_2 z + \dots + B_q z^{q-1},$$

$$(2.5) \quad C(z) = I + C_1 z + \dots + C_r z^r,$$

$$(2.6) \quad \theta^\tau = [-A_1 \dots -A_p B_1 \dots B_q C_1 \dots C_r].$$

Denote by  $\theta_n$  the  $n$ th estimate for  $\theta$ , and let  $\theta_n$  be given by

$$(2.7) \quad \theta_{n+1} = \theta_n + \frac{\varphi_n}{r_n} (y_{n+1}^\tau - \varphi_n^\tau \theta_n)$$

with

$$(2.8) \quad \varphi_n^\tau = [y_n^\tau, y_{n-1}^\tau, \dots, y_{n-p+1}^\tau, u_n^\tau \dots u_{n-q+1}^\tau, y_n^\tau - \varphi_{n-1}^\tau \theta_{n-1}, \dots, y_{n-r+1}^\tau - \varphi_{n-r}^\tau \theta_{n-r}],$$

$$(2.9) \quad r_n = 1 + \sum_{i=1}^n \|\varphi_i\|^2, \quad r_0 = 1.$$

The initial values  $\theta_0$  and  $\varphi_0$  are arbitrarily chosen.

Under reasonable conditions Goodwin, Ramadge and Caines [15] proved the global convergence and asymptotical optimality of the tracking system with  $u_n$  defined from

$$(2.10) \quad \theta_n^\tau \varphi_n = y_{n+1}^*$$

i.e.,

$$(2.11) \quad \overline{\lim}_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n \|u_i\|^2 < \infty, \quad \overline{\lim}_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n \|y_i\|^2 < \infty \quad \text{a.s.}$$

$$(2.12) \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n (y_i - y_i^*)(y_i - y_i^*)^\tau = R \quad \text{a.s.}$$

However, in this case the estimate  $\theta_n$  may be inconsistent (Becker, Kumar and Wei [2]). It can be easily explained by the following example. Let  $y_n^* \equiv 0$  and  $\theta_0^\tau \theta_0 > \theta^\tau \theta$ . Then we have

$$\theta_n^\tau \varphi_n \equiv 0, \quad \theta_n^\tau (\theta_{n+1} - \theta_n) = \theta_n^\tau \frac{\varphi_n}{r_n} y_{n+1}^\tau \equiv 0;$$

hence

$$\begin{aligned} \theta_n^\tau \theta_n &= \theta_{n-1}^\tau \theta_{n-1} + (\theta_n - \theta_{n-1})^\tau (\theta_n - \theta_{n-1}) \\ &= \theta_0^\tau \theta_0 + \sum_{i=1}^n (\theta_i - \theta_{i-1})^\tau (\theta_i - \theta_{i-1}) \geq \theta_0^\tau \theta_0 > \theta^\tau \theta. \end{aligned}$$

In order to achieve strongly consistent parameter estimates, Caines and Lafortune [3], Chen [7] and Chen and Caines [9] added a disturbance with covariance matrix  $R_1 > 0$  to the reference sequence  $\{y_n^*\}$ . In this case  $\theta_n$  tends to  $\theta$  but the long run average of the tracking errors differs from its minimum value  $R$  by an additional term  $R_1$ :

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n (y_i - y_i^*)(y_i - y_i^*)^\tau = R + R_1 \quad \text{a.s.}$$

It is natural to ask: Is it possible to achieve simultaneously both asymptotic optimality of the adaptive tracking system and the strong consistency of parameter estimates? To answer this question is the topic of the paper.

**3. Main result.** We first define adaptive control for the tracking system. Let  $\{\varepsilon_i\}$  be an  $m$ -dimensional i.i.d. sequence which is independent of  $\{w_n\}$  with properties  $E\varepsilon_i \varepsilon_i^\tau = I, E\|\varepsilon_i\|^5 < \infty$ .

Without loss of generality we assume  $\mathcal{F}_n = \sigma\{w_i, i \leq n; \varepsilon_j, j \leq n\}$ .

Unlike (2.10) we define adaptive control from the equation

$$(3.1) \quad \theta_n^\tau \varphi_n = y_{n+1}^* + v_n$$

where  $\{y_n^*\}$  is a bounded deterministic reference sequence and

$$(3.2) \quad v_1 = 0, \quad v_n = \frac{\varepsilon_n}{\log^{1/8} n} \quad \forall n \geq 2.$$

(The existence of  $u_n$  satisfying (3.1) or (2.10) is discussed in Appendix 1.)

The disturbance  $v_n$  in (3.1) is designed to have a vanishing covariance matrix in order to make tracking error asymptotically minimal, but for this the system loses the persistent excitation property which is of crucial importance in the analysis of Caines and Lafortune [3], Chen [7] and Chen and Caines [9]. To overcome this difficulty is the main task of the present paper.

We need the following conditions:

- (A<sub>1</sub>)  $C(z) - \frac{1}{2}I$  is strictly positive real;
- (A<sub>2</sub>)  $B_1$  if of full rank and zeros of  $\det B_1^+ B(z)$  lie outside the closed unit disk;
- (A<sub>3</sub>)  $B_1^+ A(z)$  and  $B_1^+ B(z)$  are left-coprime and  $B_1^+ B_q$  is of full rank;
- (A<sub>4</sub>)  $\{w_i\}$  is a mutually independent sequence with  $Ew_i = 0; \sup_i E\|w_i\|^{4+\delta} < \infty$  for some  $\delta > 0$  and

$$(3.3) \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n w_i w_i^\tau = R > 0 \quad \text{a.s.}$$

**THEOREM 1.** For system (2.1)–(2.2) let the parameter estimate be given by (2.7)–(2.9) and let the control be defined by (3.1) and (3.2). If conditions  $A_1$ – $A_4$  are fulfilled, then the tracking system is asymptotically optimal and the estimate is strongly consistent, i.e., (2.11) and (2.12) take place and  $\theta_n \xrightarrow{n \rightarrow \infty} \theta$ . a.s.

The proof of this theorem is given in § 5. For this we need some criteria for strong consistency of parameter estimate for systems without monitoring.

**4. Parameter estimation for systems without monitoring.** In contrast to  $\varphi_n$  we define an estimate-free vector  $\varphi_n^0$ :

$$(4.1) \quad \varphi_n^0 = [y_n^\tau \cdots y_{n-p+1}^\tau, u_n^\tau \cdots u_{n-q+1}^\tau, w_n^\tau \cdots w_{n-r+1}^\tau]^\tau$$

and set

$$(4.2) \quad \xi_n = y_n - w_n - \theta_{n-1}^\tau \varphi_{n-1},$$

$$(4.3) \quad \varphi_n^\xi = [0 \cdots 0, 0 \cdots 0, \xi_n^\tau \cdots \xi_{n-r+1}^\tau]^\tau.$$

Then we have

$$(4.4) \quad \varphi_n = \varphi_n^0 + \varphi_n^\xi,$$

$$(4.5) \quad y_{n+1} = \theta^\tau \varphi_n^0 + w_{n+1},$$

and

$$\begin{aligned} \theta_{n+1} &= \theta_n + \frac{\varphi_n}{r_n} (\varphi_n^{0\tau} \theta + w_{n+1}^\tau - \varphi_n^\tau \theta_n) \\ &= \theta_n + \frac{\varphi_n}{r_n} (\varphi_n^\tau \theta - \varphi_n^{\xi\tau} \theta + w_{n+1}^\tau - \varphi_n^\tau \theta_n); \end{aligned}$$

hence

$$(4.6) \quad \tilde{\theta}_{n+1} = \left( I - \frac{\varphi_n \varphi_n^\tau}{r_n} \right) \tilde{\theta}_n + \frac{\varphi_n \varphi_n^{\xi\tau}}{r_n} \theta - \frac{\varphi_n}{r_n} w_{n+1}^\tau$$

with

$$(4.7) \quad \tilde{\theta}_n = \theta - \theta_n.$$

Let the matrix  $\Phi(n, i)$  be recursively defined by

$$(4.8) \quad \Phi(n+1, i) = \left( I - \frac{\varphi_i \varphi_i^\tau}{r_i} \right) \Phi(n, i), \quad \Phi(i, i) = I.$$

Then from (4.6) it follows that

$$(4.9) \quad \tilde{\theta}_{n+1} = \Phi(n+1, 0) \tilde{\theta}_0 + \sum_{j=0}^n \Phi(n+1, j+1) \frac{\varphi_j \varphi_j^{\xi\tau}}{r_j} \theta - \sum_{j=0}^n \Phi(n+1, j+1) \frac{\varphi_j}{r_j} w_{j+1}^\tau,$$

from which we see that the behavior of  $\Phi(n, 0)$  is of great importance for consistency of parameter estimates.

**LEMMA 1.** For the system and algorithm defined by (2.1)–(2.2) and (2.7)–(2.9) if condition  $A_1$  holds, then

$$(4.10) \quad \sum_{n=0}^{\infty} \frac{\|\xi_{n+1}\|^2}{r_n} < \infty \quad \text{a.s.};$$

moreover, if conditions  $A_2$  and  $A_4$  hold and (2.10) or (3.1) is satisfied, then  $r_n \rightarrow \infty$ , and

$$(4.11) \quad \frac{1}{n} \sum_{i=0}^n \|\xi_{i+1}\|^2 \xrightarrow{n \rightarrow \infty} 0.$$

*Proof.* In Chen and Caines [9] and Chen [7], (4.10) and (4.11) are proved for  $v_n$  with the constant covariance matrix, but they can be verified by the same argument used there.  $\square$

LEMMA 2. *For the system and algorithm defined by (2.1)–(2.2) and (2.7)–(2.9) if condition  $A_1$  holds then  $\Phi(n, 0) \xrightarrow{n \rightarrow \infty} 0$  implies  $\theta_n \xrightarrow{n \rightarrow \infty} \theta$  a.s. for any initial value  $\theta_0$ . For the special case of  $r = 0$ , the converse assertion is also true, i.e., if  $\theta_n \xrightarrow{n \rightarrow \infty} \theta$  a.s. for any  $\theta_0$  then  $\Phi(n, 0) \xrightarrow{n \rightarrow \infty} 0$  a.s.*

*Proof.* The first step is to show

$$(4.12) \quad \sum_{i=0}^{n-1} \frac{\|\Phi(n, i+1)\varphi_i\|^2}{r_i} \leq d$$

for any vector sequence  $\{\varphi_n\}$  with  $\Phi(n, i)$  and  $r_n$  related by (2.9) and (4.8), where  $d$  is the dimension of  $\varphi_n$ .

Then by (2.2) and (4.12) we can prove that the last term of (4.9) goes to zero if  $\Phi(n, 0) \xrightarrow{n \rightarrow \infty} 0$ . Finally, by (4.10) and (4.12) the second term on the right-hand side of (4.9) also converges to zero if  $\Phi(n, 0) \xrightarrow{n \rightarrow \infty} 0$ . This is just a sketch proof for the first conclusion. For detailed proof we refer to Chen and Guo [12]. The second conclusion can be easily seen from (4.9).  $\square$

LEMMA 3. *If  $r_n \xrightarrow{n \rightarrow \infty} \infty$ ,  $\overline{\lim}_{n \rightarrow \infty} r_n/r_{n-1} < \infty$  and there exist quantities  $N_0$  and  $M$  possibly depending on  $\omega$  such that*

$$(4.13) \quad \frac{\lambda_{\max}^n}{\lambda_{\min}^n} \leq M(\log r_n)^{1/4} \quad \text{a.s.} \quad \forall n \geq N_0;$$

*then  $\Phi(n, 0) \xrightarrow{n \rightarrow \infty} 0$ , where  $\lambda_{\max}^n$  and  $\lambda_{\min}^n$  denote the maximum and minimum eigenvalue of the matrix  $\sum_{i=1}^n \varphi_i \varphi_i^T + (1/d)I$  respectively and  $d$  denotes the dimension of  $\varphi_n$ .*

*Proof.* We only give a sketch of the proof and refer readers interested in details to Chen and Guo [10], [11].

The key point is to find a function  $m(t)$  such that  $m(t) \xrightarrow{t \rightarrow \infty} \infty$  and

$$\|\Phi(m(N+k\alpha), m(N+(k-1)\alpha))\| \leq \sqrt{1 - \frac{\beta^2}{c_1 k}} \quad \forall k \geq 1$$

for some  $N, \alpha > 0, \beta > 0$  and  $c_1 > 0$ . If it has been done, then

$$\begin{aligned} \|\Phi(m(N+k\alpha), 0)\| &\leq \prod_{i=1}^k \|\Phi(m(N+i\alpha), m(N+(i-1)\alpha))\| \cdot \|\Phi(m(N), 0)\| \\ &\leq \left[ \prod_{i=1}^k \left(1 - \frac{\beta^2}{c_1 i}\right) \right]^{1/2} \xrightarrow{k \rightarrow \infty} 0; \end{aligned}$$

hence  $\Phi(n, 0) \xrightarrow{n \rightarrow \infty} 0$  since  $\|\Phi(j+1, j)\| \leq 1$  for all  $j$ .

It appears that the following defined function can serve as the desired one:

$$m(t) = \max [n: t_n \leq t],$$

$$t_n \triangleq \sum_{i=2}^{n-1} \frac{\|\varphi_i\|^2}{r_i (\log r_i)^{1/4}}. \quad \square$$

*Remark 1.* Lemma 3 is a purely algebraic result, namely, it is true for any vector sequence  $\{\varphi_n\}$ , only if  $\Phi(n, 0), \varphi_n$  and  $r_n$  are related by (2.9) and (4.8).

*Remark 2.* There exists an example (see Chen and Guo [13]) showing that Lemma 3 is no longer true if condition (4.13) is replaced by a more general one:

$$\lambda_{\max}^n / \lambda_{\min}^n \leq M(\log r_n)^{1+a}, \quad a > 0.$$

This means that, in order for the estimate given by (2.7) to be consistent, the condition number of  $\sum_{i=1}^n \varphi_i \varphi_i^T + (1/d)I$  is allowed to diverge at a rate of  $(\log r_n)^{1/4}$ , but not faster than  $(\log r_n)^{1+a}$ .

LEMMA 4. Let  $\{\varphi_n^1\}$ ,  $\{\varphi_n^2\}$  and  $\{\psi_n\}$  be the vector sequence satisfying conditions  $\varphi_n^1 = \varphi_n^2 + \psi_n$  and

$$(4.14) \quad \sum_{n=0}^{\infty} \frac{\|\psi_n\|^2}{r_{1n}} < \infty.$$

Then  $\Phi_1(n, 0) \xrightarrow{n \rightarrow \infty} 0$  if and only if  $\Phi_2(n, 0) \xrightarrow{n \rightarrow \infty} 0$ , where by definition

$$\Phi_i(n+1, 0) = \left( I - \frac{\varphi_n^i \varphi_n^{i\tau}}{r_{in}} \right) \Phi_i(n, 0), \quad \Phi_i(0, 0) = I,$$

$$r_{in} = 1 + \sum_{j=1}^n \|\varphi_j^i\|^2, \quad r_{i0} = 1, \quad i = 1, 2.$$

*Proof.* Without loss of generality we assume that  $\|\varphi_0^1\| \neq 1$ .

Suppose  $\Phi_1(n, 0) \xrightarrow{n \rightarrow \infty} 0$ ; then from the following chain of equalities:

$$\begin{aligned} \det \Phi_1(n+1, 0) &= \det \prod_{i=0}^n \Phi_1(i+1, i) = \prod_{i=0}^n \det \left( I - \frac{\varphi_i^1 \varphi_i^{1\tau}}{r_{1i}} \right) \\ &= \prod_{i=1}^n \frac{r_{1i-1}}{r_{1i}} (1 - \|\varphi_0^1\|^2) = \frac{1}{r_{1n}} (1 - \|\varphi_0^1\|^2) \end{aligned}$$

we see that  $r_{1n} \xrightarrow{n \rightarrow \infty} \infty$ .

By (4.14) and the Kronecker lemma we have

$$(4.15) \quad \frac{r_{2n}}{r_{1n}} = \frac{r_{1n} - 2 \sum_{i=1}^n \varphi_i^{1\tau} \psi_i + \sum_{i=1}^n \|\psi_i\|^2}{r_{1n}} \xrightarrow{n \rightarrow \infty} 1$$

and by (4.14)

$$(4.16) \quad \sum_{n=0}^{\infty} \frac{\|\psi_n\|^2}{r_{2n}} < \infty.$$

We immediately verify that

$$\begin{aligned} \Phi_2(n+1, 0) &= \Phi_1(n+1, 0) + \sum_{j=0}^n \Phi_1(n+1, j+1) \frac{\varphi_j^1 \psi_j^T}{r_{1j}} \Phi_2(j, 0) \\ &\quad + \sum_{j=0}^n \Phi_1(n+1, j+1) \frac{\psi_j \varphi_j^{2\tau}}{r_{2j}} \Phi_2(j, 0) \\ &\quad + \sum_{j=0}^n \frac{\Phi_1(n+1, j+1) \varphi_j^1}{r_{1j}^{1/2}} \left( \sqrt{\frac{r_{2j}}{r_{1j}}} - \sqrt{\frac{r_{1j}}{r_{2j}}} \right) \frac{\varphi_j^{2\tau} \Phi_2(j, 0)}{r_{2j}^{1/2}}. \end{aligned}$$

By using (4.12) and (4.14)–(4.16) it is not difficult to conclude that  $\Phi_1(n, 0) \xrightarrow{n \rightarrow \infty} 0$  implies  $\Phi_2(n, 0) \xrightarrow{n \rightarrow \infty} 0$ . The converse implication is proved in a similar way.  $\square$

**THEOREM 2.** For the system and algorithm defined by (2.1)-(2.2) and (2.7)-(2.9) if condition  $A_1$  holds and if  $r_n \xrightarrow{n \rightarrow \infty} \infty$ ,

$$\overline{\lim}_{n \rightarrow \infty} r_n / r_{n-1} < \infty \text{ and } \lambda_{\max}^n / \lambda_{\min}^n \leq M(\log r_n)^{1/4} \quad \forall n \geq N$$

(or  $r_n^0 \xrightarrow{n \rightarrow \infty} \infty$ ,  $\overline{\lim}_{n \rightarrow \infty} r_n^0 / r_{n-1}^0 < \infty$  and  $\lambda_{\max}^{0n} / \lambda_{\min}^{0n} \leq M(\log r_n^0)^{1/4}$  for all  $n \geq N$ ) with  $N$  and  $M$  possibly depending on  $\omega$ , then

$$\theta_n \xrightarrow{n \rightarrow \infty} \theta \quad \text{a.s.}$$

for any initial value, where  $\lambda_{\max}^{0n}$ ,  $\lambda_{\min}^{0n}$  denote the maximum and minimum eigenvalue of  $\sum_{i=1}^n \varphi_i^0 \varphi_i^{0\tau} + (1/d)I$ , respectively, and  $r_n^0 = 1 + \sum_{i=1}^n \|\varphi_i^0\|^2$  with  $\varphi_n^0$  defined by (4.1).

*Proof.* Since (4.4), (4.10) and Lemma 4 can be applied with  $\varphi_n^1 = \varphi_n$ ,  $\varphi_n^2 = \varphi_n^0$  and  $\psi_n = \varphi_n^\xi$ , hence  $\Phi(n, 0) \xrightarrow{n \rightarrow \infty} 0$  if and only if  $\Phi_0(n, 0) \xrightarrow{n \rightarrow \infty} 0$ , where  $\Phi_0(n, 0)$  is defined by (4.8) with  $\varphi_n$  and  $r_n$  replaced by  $\varphi_n^0$  and  $r_n^0$ , respectively. Then the conclusions of the theorem immediately follow from Lemmas 2 and 3.  $\square$

**5. Proof of Theorem 1.** To begin with we prove the following lemmas.

**LEMMA 5.** Let  $\{v_n\}$  be defined by (3.2) and let  $H_N(z) = \sum_{i=0}^{\infty} H_i(N)z^i$  be the matrix series in shift-back operator  $z$ , where the matrix coefficients  $H_i(N)$  may depend upon  $\omega$ , but there are constants (independent of  $\omega$ )  $k_1 > 0$  and  $k_2 > 0$  such that

$$\|H_i(N)\| \leq k_1 \exp(-k_2 i) \quad \forall i \quad \forall N \geq 0.$$

Then

$$\begin{aligned} \lim_{N \rightarrow \infty} \frac{\log^{1/4} N}{N} \sum_{n=1}^N (H_N(z)v_n)(H_N(z)v_n)^\tau \\ = \lim_{N \rightarrow \infty} \frac{\log^{1/4} N}{N} \sum_{i=0}^N H_i(N)H_i^\tau(N) \left( \sum_{j=2}^N \frac{1}{\log^{1/4} j} \right). \end{aligned}$$

*Proof.* See Appendix 2.

**LEMMA 6.** Let condition  $A_4$  except (3.3) be held and let  $H(z) = \sum_{i=0}^{\infty} H_i z^i$  and  $G(z) = \sum_{i=0}^{\infty} G_i z^i$  be matrix series in shift-back operator  $z$  with  $\|H_i\| + \|G_i\| \leq k_1 \exp(-k_2 i)$  for all  $i \geq 0$ , for some constants  $k_1 > 0$ ,  $k_2 > 0$ . Then there exists  $\gamma \in (0, 1)$  such that for all  $l \geq 0$ ,  $m \geq 0$ ,

$$(5.1) \quad \lim_{N \rightarrow \infty} \frac{1}{N^\gamma} \sum_{n=1}^N (H(z)w_{n+1-l})(G(z)v_{n-m})^\tau = 0 \quad \text{a.s.},$$

$$(5.2) \quad \lim_{N \rightarrow \infty} \frac{1}{N^\gamma} \sum_{n=1}^N (H(z)w_{n+1-l})\eta_n^\tau = 0 \quad \text{a.s.}$$

for any bounded deterministic sequence  $\{\eta_n\}$ , and

$$(5.3) \quad \overline{\lim}_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N \|H(z)w_{n+1-l}\|^2 < \infty \quad \text{a.s.}$$

*Proof.* See Appendix 2.

Set

$$(5.4) \quad H_1(z) = [B_1^+ B(z)]^{-1} B_1^+ A(z),$$

$$(5.5) \quad H_2(z) = H_1(z) - [B_1^+ B(z)]^{-1} B_1^+ C(z),$$

$$(5.6) \quad Y_n^* = [y_n^{*\tau} \cdots y_{n-p+1}^{*\tau}, (H_1(z)y_{n+1}^*)^\tau \cdots (H_1(z)y_{n-q+2}^*)^\tau]^\tau$$

and

$$(5.7) \quad Z_n = [v_{n-1}^\tau \cdots v_{n-p}^\tau, (H_1(z)v_n)^\tau \cdots (H_1(z)v_{n-q+1})^\tau]^\tau.$$

In the following, by  $\lambda_{\min}(X)$  ( $\lambda_{\max}(X)$ ) we mean the minimum (maximum) eigenvalue of the matrix  $X$ ; we have the following.

LEMMA 7. Under conditions of Theorem 1 if

$$(5.8) \quad \lim_{N \rightarrow \infty} \lambda_{\min} \left( \frac{\log^{1/4} N}{N} \sum_{n=1}^N (Y_n^* Y_n^{*\tau} + Z_n Z_n^\tau) \right) \neq 0 \quad \text{a.s.}$$

then

$$\theta_n \xrightarrow[n \rightarrow \infty]{} \theta \quad \text{a.s.}$$

*Proof.* By Theorem 2 we only need to prove  $\Phi_0(n, 0) \xrightarrow[n \rightarrow \infty]{} 0$  a.s. From (2.1) we have

$$(5.9) \quad u_n = [B_1^+ B(z)]^{-1} B_1^+ A(z) y_{n+1} - [B_1^+ B(z)]^{-1} B_1^+ C(z) w_{n+1};$$

then by (3.1), (4.2), (5.4), (5.5) and (5.9),  $\varphi_n^0$  defined by (4.1) can be written as

$$(5.10) \quad \varphi_n^0 = \varphi_n^1 + \psi_n$$

where

$$(5.11) \quad \psi_n = [\xi_n^\tau, \cdots, \xi_{n-p+1}^\tau, (H_1(z)\xi_{n+1})^\tau \cdots (H_1(z)\xi_{n-q+2})^\tau, 0 \cdots 0]^\tau,$$

$$(5.12) \quad \varphi_n^1 = \varphi_n^2 + \varphi_n^3,$$

$$(5.13) \quad \varphi_n^2 = [w_n^\tau \cdots w_{n-p+1}^\tau, (H_2(z)w_{n+1})^\tau \cdots (H_2(z)w_{n-q+2})^\tau, w_n^\tau \cdots w_{n-r+1}^\tau]^\tau,$$

$$(5.14) \quad \varphi_n^3 = [Y_n^{*\tau} + Z_n^\tau, 0 \cdots 0]^\tau.$$

By (4.4), (4.10), similar to (4.15) we have

$$(5.15) \quad \frac{r_n^0}{r_n} = \frac{r_n - 2 \sum_{i=1}^n \varphi_i^\tau \varphi_i^\xi + \sum_{i=1}^n \|\varphi_i^\xi\|^2}{r_n} \xrightarrow[n \rightarrow \infty]{} 1.$$

Then by the Schwarz inequality it follows that

$$\begin{aligned} \sum_{n=0}^{\infty} \frac{\|H_1(z)\xi_{n+1-l}\|^2}{r_n^0} &\leq \sum_{n=0}^{\infty} \frac{1}{r_n^0} \sum_{i=0}^{\infty} \|H_{1i}\| \sum_{i=0}^{\infty} \|H_{1i}\| \cdot \|\xi_{n+1-l-i}\|^2 \\ &\leq k_0 \sum_{i=0}^{\infty} \|H_{1i}\| \sum_{n=0}^{\infty} \frac{\|\xi_{n+1-l-i}\|^2}{r_n^0} < \infty \end{aligned}$$

where the last inequality is obtained because  $\xi_i = 0$  for  $i < 0$  and the coefficients in  $H_1(z) = \sum_{i=0}^{\infty} H_{1i} z^i$  have the estimates  $\|H_{1i}\| \leq k_1 \exp(-k_2 i)$ , for all  $i \geq 0$ , ( $k_1 > 0, k_2 > 0$ ) by condition  $A_2$ . Thus we have established

$$(5.16) \quad \sum_{n=0}^{\infty} \frac{\|\psi_n\|^2}{r_n^0} < \infty$$

and by Lemma 4 we conclude that  $\Phi_0(n, 0) \xrightarrow[n \rightarrow \infty]{} 0$  iff  $\Phi_1(n, 0) \xrightarrow[n \rightarrow \infty]{} 0$ .

Next, we prove that

$$(5.17) \quad \lim_{N \rightarrow \infty} \lambda_{\min} \left( \frac{\log^{1/4} N}{N} \sum_{n=1}^N \varphi_n^1 \varphi_n^{1\tau} \right) \neq 0 \quad \text{a.s.}$$

If for some  $\omega \in \Omega$  (5.17) were not true, then we would find a subsequence of eigenvectors  $\begin{bmatrix} \alpha_{N_k} \\ \beta_{N_k} \end{bmatrix}$  for matrix  $(\log^{1/4} N_k / N_k) \sum_{n=1}^{N_k} \varphi_n^1 \varphi_n^{1\tau}$  with  $N_k \xrightarrow[k \rightarrow \infty]{} \infty$ ,  $\alpha_{N_k} \in R^{mp+1q}$ ,  $\beta_{N_k} \in R^{mr}$  and

$$(5.18) \quad \|\alpha_{N_k}\|^2 + \|\beta_{N_k}\|^2 = 1$$

such that

$$(5.19) \quad (\alpha_{N_k}^\tau, \beta_{N_k}^\tau) \frac{\log^{1/4} N_k}{N_k} \sum_{n=1}^{N_k} \varphi_n^1 \varphi_n^{1\tau} \begin{pmatrix} \alpha_{N_k} \\ \beta_{N_k} \end{pmatrix} \xrightarrow[k \rightarrow \infty]{} 0.$$

Without loss of generality we always assume that this fixed  $\omega$  does not belong to a possible exceptional set of probability zero. Obviously,  $\alpha_{N_k}$  and  $\beta_{N_k}$  would be  $\omega$ -dependent but not necessarily measurable.

Utilizing Lemma 6 one can easily be convinced of the fact

$$(5.20) \quad \frac{\log^{1/4} N}{N} \sum_{n=1}^N \varphi_n^2 \varphi_n^{3\tau} \xrightarrow[N \rightarrow \infty]{} 0,$$

then (5.19) is reduced to

$$(5.21) \quad (\alpha_{N_k}^\tau, \beta_{N_k}^\tau) \frac{\log^{1/4} N_k}{N_k} \sum_{n=1}^{N_k} \varphi_n^2 \varphi_n^{2\tau} \begin{pmatrix} \alpha_{N_k} \\ \beta_{N_k} \end{pmatrix} \xrightarrow[k \rightarrow \infty]{} 0,$$

$$(5.22) \quad (\alpha_{N_k}^\tau, \beta_{N_k}^\tau) \frac{\log^{1/4} N_k}{N_k} \sum_{n=1}^{N_k} \varphi_n^3 \varphi_n^{3\tau} \begin{pmatrix} \alpha_{N_k} \\ \beta_{N_k} \end{pmatrix} \xrightarrow[k \rightarrow \infty]{} 0.$$

In view of Lemma 6, (5.8) implies

$$(5.23) \quad \varliminf_{N \rightarrow \infty} \lambda_{\min} \left( \frac{\log^{1/4} N}{N} \sum_{n=1}^N (Y_n^* + Z_n)(Y_n^* + Z_n)^\tau \right) \neq 0.$$

Paying attention to the fact that the last  $mr$  elements in  $\varphi_n^3$  are zeros, by (5.22) and (5.23) we conclude that

$$(5.24) \quad \alpha_{N_k} \xrightarrow[k \rightarrow \infty]{} 0;$$

hence, recalling (5.18) we have

$$(5.25) \quad \|\beta_{N_k}\| \xrightarrow[k \rightarrow \infty]{} 1.$$

Let

$$x_n^1 = [w_n^\tau \cdots w_{n-p+1}^\tau, (H_2(z)w_{n+1})^\tau \cdots (H_2(z)w_{n-q+2})^\tau]^\tau,$$

$$x_n^2 = [w_n^\tau \cdots w_{n-r+1}^\tau]^\tau.$$

Then  $\varphi_n^2 = [x_n^1, x_n^2]^\tau$  and (5.21) implies

$$(5.26) \quad \frac{1}{N_k} \sum_{n=1}^{N_k} \|\alpha_{N_k}^\tau x_n^1 + \beta_{N_k}^\tau x_n^2\|^2 \xrightarrow[k \rightarrow \infty]{} 0.$$

Further, we have

$$(5.27) \quad \varliminf_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N \|x_n^1\|^2 < \infty \quad \text{a.s.}$$

by Lemma 6, and

$$(5.28) \quad \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N x_n^2 x_n^{2\tau} = \begin{bmatrix} R & 0 \\ 0 & R \end{bmatrix} > 0$$

by ergodicity.

Thus from (5.24) and (5.26)–(5.28) it follows that

$$\lim_{k \rightarrow \infty} \frac{1}{N_k} \sum_{n=1}^{N_k} \beta_{N_k}^\tau x_n^2 x_n^{2\tau} \beta_{N_k} = 0,$$

which leads to  $\beta_{N_k} \xrightarrow{k \rightarrow \infty} 0$  by (5.28). Comparing it with (5.25) we obtain a contradiction, which shows the truth of (5.17). Therefore, there exist  $\alpha_0 > 0, N_0$  such that

$$(5.29) \quad \lambda_{\min} \left( \sum_{i=1}^n \varphi_i^1 \varphi_i^{1\tau} \right) \geq \frac{n}{\log^{1/4} n} \alpha_0 \quad \forall n \geq N_0.$$

By (5.12) and Lemma 6 it follows that

$$(5.30) \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n \|\varphi_i^1\|^2 \geq \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n \|w_i\|^2 = \text{tr } R > 0$$

and

$$(5.31) \quad \overline{\lim}_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n \|\varphi_i^1\|^2 < \infty \quad \text{a.s.}$$

From (5.30) and (5.31) it follows that there are positive quantities  $\beta \geq \alpha > 0$  such that

$$(5.32) \quad \alpha n \leq r_{1n} \leq \beta n,$$

which, together with (5.29), yields

$$\lambda_{\max}^{1n} / \lambda_{\min}^{1n} \leq M \log^{1/4} r_{1n} \quad \forall n \geq N_0 \quad \text{with some } M > 0$$

where  $\lambda_{\max}^{1n}$  and  $\lambda_{\min}^{1n}$  denote, respectively, the maximum and minimum eigenvalues of  $\sum_{i=1}^n \varphi_i^1 \varphi_i^{1\tau} + (1/\alpha)I$ . Then we obtain the required assertion  $\Phi_1(n, 0) \xrightarrow{n \rightarrow \infty} 0$  by Lemma 3 and Remark 1.  $\square$

*Proof of Theorem 1.* Since  $\sum_{i=2}^n (v_i v_i^\tau - (1/\log^{1/4} i)I)/i$  is a convergent martingale, by the Kronecker lemma it follows that

$$(5.33) \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n v_i v_i^\tau = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=2}^n \frac{1}{\log^{1/4} i} I = 0.$$

Similarly we have

$$(5.34) \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n w_i v_{i-1}^\tau = 0.$$

From (3.1) and (4.2) it follows that

$$(5.35) \quad y_{n+1} = \xi_{n+1} + y_{n+1}^* + w_{n+1} + v_n.$$

Then (2.11) and (2.12) follow immediately from (4.11), (5.9), (5.33)–(5.35) and condition  $A_2$ .

Thus we only need to prove  $\theta_n \xrightarrow{n \rightarrow \infty} \theta$  a.s. By Lemma 7 it suffices to verify that

$$(5.36) \quad \lim_{N \rightarrow \infty} \lambda_{\min} \left( \frac{\log^{1/4} N}{N} \sum_{n=1}^N Z_n Z_n^\tau \right) \neq 0 \quad \text{a.s.}$$

If (5.36) were not true, then there would exist a subsequence of eigenvectors  $\alpha_{N_k} \in R^{mp+lq}$  for matrix  $(\log^{1/4} N_k / N_k) \sum_{n=1}^{N_k} Z_n Z_n^T$  with  $N_k \xrightarrow[k \rightarrow \infty]{} \infty$  and

$$(5.37) \quad \|\alpha_{N_k}\| = 1 \quad \forall k \geq 1$$

such that

$$(5.38) \quad \alpha_{N_k}^T \frac{\log^{1/4} N_k}{N_k} \sum_{n=1}^{N_k} Z_n Z_n^T \alpha_{N_k} \xrightarrow[k \rightarrow \infty]{} 0.$$

Without loss of generality we suppose  $\alpha_{N_k} \xrightarrow[k \rightarrow \infty]{} \alpha$ . Write  $\alpha_{N_k}$  and  $\alpha$  in the component form

$$\alpha_{N_k} = [\alpha_1^T(N_k) \cdots \alpha_{p+q}^T(N_k)]^T, \quad \alpha = [\alpha_1^T \cdots \alpha_{p+q}^T]^T$$

with  $\alpha_i(N_k)$ ,  $\alpha_i$  being  $m$ -dimensional and  $\alpha_{p+j}(N_k)$ ,  $\alpha_{p+j}$   $l$ -dimensional vectors,  $i = 1 \cdots p, j = 1 \cdots q$ .

Set

$$(5.39) \quad \begin{aligned} H_{N_k}(z) &= \alpha_1^T(N_k)z + \cdots + \alpha_p^T(N_k)z^p \\ &+ \alpha_{p+1}^T(N_k)H_1(z) + \cdots + \alpha_{p+q}^T(N_k)H_1(z)z^{q-1} \\ &\triangleq \sum_{i=0}^{\infty} h_i^T(N_k)z^i, \end{aligned}$$

$$(5.40) \quad \begin{aligned} H(z) &= \alpha_1^T z + \cdots + \alpha_p^T z^p + \alpha_{p+1}^T H_1(z) + \cdots + \alpha_{p+q}^T H_1(z)z^{q-1} \\ &\triangleq \sum_{i=0}^{\infty} h_i^T z^i. \end{aligned}$$

We note that  $\alpha_{N_k}$  and hence  $H_{N_k}(z)$  may depend on  $\omega$ , but by condition  $A_2$  and (5.37) it is clear that there are constants  $c_1 > 0, c_2 > 0$  such that  $\|h_i(N_k)\| \leq c_1 \exp(-c_2 i)$  for all  $i \geq 0$ , for all  $k \geq 0$ . Then Lemma 5 can be applied, and from (5.38), (5.39) we have

$$\begin{aligned} 0 &= \lim_{k \rightarrow \infty} \frac{\log^{1/4} N_k}{N_k} \sum_{n=1}^{N_k} [\alpha_1^T(N_k)v_{n-1} + \cdots + \alpha_p^T(N_k)v_{n-p} + \alpha_{p+1}^T(N_k)H_1(z)v_n + \cdots \\ &\quad + \alpha_{p+q}^T(N_k)H_1(z)v_{n-q+1}]^2 \\ &= \lim_{k \rightarrow \infty} \frac{\log^{1/4} N_k}{N_k} \sum_{n=1}^{N_k} [(\alpha_1^T(N_k)z + \cdots + \alpha_p^T(N_k)z^p + \alpha_{p+1}^T(N_k)H_1(z) + \cdots \\ &\quad + \alpha_{p+q}^T(N_k)H_1(z)z^{q-1})v_n]^2 \\ &= \lim_{k \rightarrow \infty} \frac{\log^{1/4} N_k}{N_k} \sum_{n=1}^{N_k} (H_{N_k}(z)v_n)(H_{N_k}(z)v_n)^T \\ &= \lim_{k \rightarrow \infty} \frac{\log^{1/4} N_k}{N_k} \sum_{i=0}^{N_k} h_i^T(N_k) \sum_{n=2}^{N_k} \frac{1}{\log^{1/4} n} h_i(N_k), \end{aligned}$$

which implies that

$$\lim_{k \rightarrow \infty} \sum_{i=0}^{N_k} \|h_i(N_k)\|^2 = 0$$

and hence  $\sum_{i=0}^{\infty} \|h_i\|^2 = 0$  by the dominated convergence theorem; therefore  $H(z) = 0$ . Setting  $z = 0$  and paying attention to the fact that  $H_1(0) (= B_1^+)$  is of full row rank we see  $\alpha_{p+1}^T B_1^+ = 0$  and so  $\alpha_{p+1} = 0$ . Then it follows directly from (5.40) that

$$(5.41) \quad (\alpha_1^T + \alpha_2^T z + \cdots + \alpha_p^T z^{p-1}) = -(\alpha_{p+2}^T + \cdots + \alpha_{p+q}^T z^{q-2}) H_1(z).$$

In view of condition  $A_3$ , applying Lemma 6.6-1 of Kailath [16] to (5.41) we know that there exists a polynomial with vector coefficients  $f(z)$ :

$$f(z) = f_1 + f_2z + \dots + f_s z^{s-1}, \quad s \geq 1,$$

such that

$$(5.42) \quad (\alpha_{p+2}^\tau + \dots + \alpha_{p+q}^\tau z^{q-2}) = f^\tau(z) B_1^+ B(z) \\ = (f_1^\tau + \dots + f_s^\tau z^{s-1})(B_1^+ B_1 + \dots + B_1^+ Bqz^{q-1}).$$

From here it is easy to conclude that  $f_i = 0$  ( $1 \leq i \leq s$ ) since  $B_1^+ B_q$  is of full rank by condition  $A_3$ , then  $\alpha_{p+j} = 0$  by (5.42), and then  $\alpha_j = 0$  by (5.41) ( $1 \leq j \leq q, 1 \leq i \leq p$ ). Thus  $\alpha = 0$ , and  $\alpha_{N_k} \xrightarrow{k \rightarrow \infty} 0$ , thus contradicting (5.37). Hence (5.36) holds.  $\square$

**6. Concluding discussion.** In order to get optimality in both tracking and estimating we have added to  $\{y_n^*\}$  a random disturbance with covariance matrix tending to zero, but, intuitively, the disturbance may harm the tracking if time is bounded. However, all assertions of Theorem 1 can remain valid for  $u_n$  defined from (3.1) with  $v_n$  deleted (i.e. from (2.10)) if the reference signal  $y_n^*$  itself is ‘‘complicated’’ enough in the sense that

$$(6.1) \quad \lim_{N \rightarrow \infty} \lambda_{\min} \left( \frac{\log^{1/4} N}{N} \sum_{n=1}^N Y_n^* Y_n^{*\tau} \right) \neq 0.$$

This remark can easily be seen from Lemma 7.

For the single-input and single-output system it is easy to show that for (6.1) it suffices to require condition  $A_3$  and

$$(6.2) \quad \lim_{N \rightarrow \infty} \lambda_{\min} \left( \frac{\log^{1/4} N}{N} \sum_{n=1}^N [y_n^* \dots y_{n-p-q+1}^*]^T [y_n^* \dots y_{n-p-q+1}^*] \right) \neq 0.$$

Recently, for multidimensional and random  $\{y_n^*\}$  we have obtained conditions similar to (6.2) in order that all conclusions of Theorem 1 hold by applying  $u_n$  defined from (2.10). It will be published elsewhere.

**Appendix 1. Existence of adaptive control.**

LEMMA. (1) Let  $A$  and  $B$  be two matrices of dimensions  $m \times n$  and  $n \times m$ , respectively. Then the following equality takes place

$$\det(I_m + AB) = \det(I_n + BA)$$

where  $I_n$  means the  $n \times n$  identity matrix.

(2) Provided  $x_1$  and  $x_2$  are independent random variables, then

$$\sup_{a \in R^1} P(x_1 + x_2 = a) \leq \min \left\{ \sup_{a \in R^1} P(x_1 = a), \sup_{a \in R^1} P(x_2 = a) \right\}.$$

Proof. (1) By taking determinants for both sides of the following matrix identity:

$$\begin{bmatrix} I_m & -A \\ 0 & BA + I_n \end{bmatrix} = \begin{bmatrix} I_m & 0 \\ -B & I_n \end{bmatrix} \cdot \begin{bmatrix} I_m + AB & -A \\ 0 & I_n \end{bmatrix} \cdot \begin{bmatrix} I_m & 0 \\ B & I_n \end{bmatrix},$$

the desired equality is immediately verified.

(2) Denote by  $F_1(x), F_2(x), F_{12}(x)$  the distributions of  $x_1, x_2, x_1 + x_2$ , respectively. Clearly we have

$$F_{12}(x) = \int_{-\infty}^{\infty} F_1(x - y) dF_2(y)$$

and

$$F_{12}(x+) = \int_{-\infty}^{\infty} F_1((x-y)+) dF_2(y)$$

by the dominated convergence theorem.

Then for any  $a \in R^1$

$$\begin{aligned} P(x_1 + x_2 = a) &= F_{12}(a+) - F_{12}(a) = \int_{-\infty}^{\infty} [F_1((a-y)+) - F_1(a-y)] dF_2(y) \\ &= \int_{-\infty}^{\infty} P(x_1 = a-y) dF_2(y) \leq \sup_{a \in R^1} P(x_1 = a) \int_{-\infty}^{\infty} dF_2(y) \\ &= \sup_{a \in R^1} P(x_1 = a). \end{aligned}$$

Similarly, we have

$$P(x_1 + x_2 = a) \leq \sup_{a \in R^1} P(x_2 = a)$$

and thus the desired result follows.  $\square$

**THEOREM.** Assume  $m \leq l$  and  $\{w_n\}$  and  $\{v_n\}$  are two sequences of mutually independent random vectors and the components of  $w_n$  are independent and with continuous distribution functions. Then for any  $n \geq 1$  there exists  $u_n$  satisfying (3.1) if the initial values are appropriately chosen. Further, this  $u_n$  is unique if and only if  $m = l$ .

*Proof.* Let  $A_{in}, B_{jn}, C_{kn}, i = 1 \cdots p, j = 1 \cdots q, k = 1 \cdots r$  be the matrix components of  $\theta_n$ , i.e.,

$$\theta_n^\tau = [-A_{1n} \cdots -A_{pn} B_{1n} \cdots B_{qn} C_{1n} \cdots C_{rn}].$$

Set

$$\bar{\theta}_n^\tau = [-A_{1n} \cdots -A_{pn} 0 B_{2n} \cdots B_{qn} C_{1n} \cdots C_{rn}],$$

and

$$\bar{\varphi}_n = [y_n^\tau \cdots y_{n-p+1}^\tau, 0, u_{n-1}^\tau \cdots u_{n-q+1}^\tau, y_n^\tau - \varphi_{n-1}^\tau \theta_{n-1}, \cdots, y_{n-r+1}^\tau - \varphi_{n-r}^\tau \theta_{n-r}]^\tau.$$

Equation (3.1) is equivalent to

$$(A.1) \quad B_{1n} u_n = y_{n+1}^* + v_n - \bar{\theta}_n^\tau \bar{\varphi}_n.$$

First let  $m = l$ . For this case we only need to prove that  $B_{1n}$  is invertible a.s. In fact, if this is true, then from (A.1)  $u_n$  is uniquely defined by  $u_n = B_{1n}^{-1}(y_{n+1}^* + v_n - \bar{\theta}_n^\tau \bar{\varphi}_n)$ , which obviously is  $\mathcal{F}_n$ -measurable. (In adaptive tracking cases we take  $\mathcal{F}_n \triangleq \sigma\{w_i, v_i, i \leq n\}$ .)

From (2.7) and (4.2) we obtain

$$(A.2) \quad B_{1n+1} = B_{1n} + \frac{1}{r_n} (\xi_{n+1} + w_{n+1}) u_n^\tau.$$

It is easy to take initial values  $\varphi_0, \theta_0$  such that  $B_{11}$  is invertible; for example, take  $u_0 = 0$  and  $B_{10}$  invertible.

We now inductively prove that  $B_{1n}$  is nondegenerate for any  $n \geq 0$ . Assuming  $B_{1n}$  is nonsingular a.s., we show that  $B_{1n+1}$  is also. In other words, we need to prove that  $P(N) = 0$  implies  $P(DN^c) = 0$ , where

$$N \triangleq \{\omega \mid \det B_{1n} = 0\}, \quad D \triangleq \{\omega \mid \det B_{1n+1} = 0\}.$$

Suppose that the opposite were true, i.e.,  $P(N) = 0$ , but  $P(DN^c) > 0$ .

From (A.2) we have

$$\det \left( B_{1n} + \frac{1}{r_n} (\xi_{n+1} + w_{n+1}) u_n^T \right) = 0 \quad \forall \omega \in DN^c$$

but  $\det B_{1n} \neq 0$  for  $\omega \in DN^c$ ; hence

$$\det \left( I + \frac{1}{r_n} B_{1n}^{-1} (\xi_{n+1} + w_{n+1}) u_n^T \right) = 0 \quad \forall \omega \in DN^c$$

or

$$\det \left( 1 + \frac{1}{r_n} u_n^T B_{1n}^{-1} (\xi_{n+1} + w_{n+1}) \right) = 0 \quad \forall \omega \in DN^c$$

by part (1) of the lemma.

Then we have

$$(A.3) \quad u_n^T B_{1n}^{-1} (\xi_{n+1} + w_{n+1}) = -r_n \quad \forall \omega \in DN^c$$

and consequently,

$$(A.4) \quad u_n^T B_{1n} \neq 0 \quad \forall \omega \in DN^c$$

since  $r_n \geq 1$ .

We denote by  $\alpha_i(\omega)$  and  $w_{n+1,i}$  the components of  $u_n^T B_{1n}^{-1}$  and  $w_{n+1}$  respectively, i.e.,

$$(A.5) \quad u_n^T B_{1n}^{-1} = [\alpha_1(\omega), \dots, \alpha_m(\omega)],$$

$$(A.6) \quad w_{n+1} = [w_{n+1,1}, \dots, w_{n+1,m}]^T.$$

Then from (A.3), (A.5) and (A.6) we have

$$(A.7) \quad \sum_{i=1}^m \alpha_i(\omega) w_{n+1,i} + r_n + u_n^T B_{1n} \xi_{n+1} = 0 \quad \forall \omega \in DN^c.$$

From (A.4) and the assumption  $P(DN^c) > 0$  we would have some  $\alpha_i(\omega)$  and a subset  $D_1 \subset DN^c$  such that

$$(A.8) \quad \alpha_i(\omega) \neq 0 \quad \forall \omega \in D_1, \quad P(D_1) > 0.$$

Without loss of generality, we assume  $i = 1$ , and define the random variable  $z(\omega)$ :

$$z(\omega) = \begin{cases} \frac{1}{\alpha_1(\omega)} \left[ \sum_{i=2}^m \alpha_i(\omega) w_{n+1,i} + r_n + u_n^T B_{1n}^{-1} \xi_{n+1} \right], & \omega \in D_1, \\ 0, & \omega \in D_1^c, \end{cases}$$

which is clearly independent of  $w_{n+1,1}$ . By part (2) of the lemma, it follows that

$$(A.9) \quad P(w_{n+1,1} + z(\omega) = 0) = 0.$$

However, (A.7) and (A.8) would yield

$$(A.10) \quad P(w_{n+1,1} + z(\omega) = 0) \geq P(D_1) > 0.$$

The contradiction obtained proves  $P(DN^c) = 0$ , and hence the nonsingularity of  $B_{1,n+1}$  a.s.

Now assume  $m < l$ .

Let

$$B_{1n} \triangleq [\overbrace{B_{1n}^1}^m, \overbrace{B_{1n}^2}^{l-m}]m, \quad u_n^\tau \triangleq [\overbrace{u_n^{1\tau}}^m, \overbrace{u_n^{2\tau}}^{l-m}].$$

From (A.2) we see

$$B_{1n+1}^1 = B_{1n}^1 + \frac{1}{r_n}(\xi_{n+1} + w_{n+1})u_n^{1\tau}.$$

By an argument similar to that given for the  $m = l$  case we can prove that  $B_{1n}^1$  is invertible a.s. for any  $n \geq 1$  if  $\varphi_0, \theta_0$  are adequately chosen. Then (A.1) is equivalent to

$$(A.11) \quad [I, (B_{1n}^1)^{-1}B_{1n}^2]u_n = (B_{1n}^1)^{-1}(y_{n+1}^* + v_n - \bar{\theta}_n^\tau \bar{\varphi}_n)$$

or

$$u_n^1 + (B_{1n}^1)^{-1}B_{1n}^2u_n^2 = (B_{1n}^1)^{-1}(y_{n+1}^* + v_n - \bar{\theta}_n^\tau \bar{\varphi}_n).$$

Obviously, the solution of (A.11) can be expressed by

$$u_n = \begin{bmatrix} (B_{1n}^1)^{-1}(y_{n+1}^* + v_n - \bar{\theta}_n^\tau \bar{\varphi}_n - B_{1n}^2u_n^2) \\ u_n^2 \end{bmatrix} \quad \text{a.s.}$$

with any  $(l - m)$ -dimensional and  $\mathcal{F}_n$ -measurable  $u_n^2$ . This means that for the case  $m < l$  the control  $u_n$  satisfying (3.1) exists but it is not unique.  $\square$

*Remark.* Recently Caines and Meyn [4] also have shown the existence of  $u_n$  satisfying (2.10) for a one-dimensional case but under conditions different from those imposed here.

**Appendix 2. Proof of lemmas.**

*Proof of Lemma 5.* Due to the assumption  $v_n = 0$  for  $n < 0$ , we have

$$\begin{aligned} \sum_{h=1}^N (H_N(z)v_n)(H_N(z)v_n)^\tau &= \sum_{i,j=0}^\infty H_i(N) \left( \sum_{n=1}^N v_{n-i}v_{n-j}^\tau \right) H_j^\tau(N) \\ &= \sum_{i,j=0}^\infty H_i(N) \left( \sum_{n=\max(i,j,1)}^N v_{n-i}v_{n-j}^\tau \right) H_j^\tau(N). \end{aligned}$$

Set

$$S_N(i, j) = \sum_{n=\max(i,j,1)}^N [v_{n-i}v_{n-j}^\tau - \delta_{ij}R_{n-i}],$$

$$R_n = Ev_nv_n^\tau = \begin{cases} \frac{1}{\log^{1/4} n} I, & n > 1, \\ 0, & n \leq 1. \end{cases}$$

Clearly,  $S_N(i, j)$  is a martingale and by Burkholder inequality (Chow and Teicher [14]),  $C_r$ -inequality and Schwarz inequality we have

$$\begin{aligned} E \|S_N(i, j)\|^{2+\delta/2} &\leq c_1 E \left( \sum_{n=\max(i,j,1)}^N \|v_{n-i}v_{n-j}^\tau - \delta_{ij}R_{n-i}\|^2 \right)^{1+(\delta/4)} \\ &\leq c_1 N^{\delta/4} E \sum_{n=\max(i,j,1)}^N \|v_{n-i}v_{n-j}^\tau - \delta_{ij}R_{n-i}\|^{2+(\delta/2)} \\ &\leq c_2 N^{1+\delta/4} \quad \text{for any } i \geq 0, j \geq 0 \text{ and some } c_1 > 0, c_2 > 0. \end{aligned}$$

From here and the Hölder inequality it follows that for any  $\varepsilon > 0$  and

$$\gamma \in \left( \frac{2 + (\delta/4)}{2 + (\delta/2)}, 1 \right),$$

$$\begin{aligned} & P \left\{ \sum_{i,j=0}^{\infty} e^{-k_2(i+j)} \|S_N(i, j)\| > N^\gamma \cdot \varepsilon \right\} \\ & \leq \frac{1}{\varepsilon^{2+(\delta/2)} N^{\gamma(2+(\delta/2))}} E \left( \sum_{i,j=0}^{\infty} e^{-k_2(i+j)} \|S_N(i, j)\| \right)^{2+(\delta/2)} \\ & \leq c_3 \frac{1}{N^{\gamma(2+(\delta/2))}} E \sum_{i,j=0}^{\infty} (e^{-k_2(i+j)})^{1+(\delta/4)} \|S_N(i, j)\|^{2+(\delta/2)} \\ & \leq c_4 \cdot \frac{1}{N^{\gamma(2+(\delta/2)) - (1+(\delta/4))}} \quad \text{for any } N \geq 1 \text{ and some constants } c_3 > 0, c_4 > 0. \end{aligned}$$

Then by the Borel–Cantelli lemma we see

$$(A.12) \quad \overline{\lim}_{N \rightarrow \infty} \frac{1}{N^\gamma} \left\| \sum_{i,j=0}^{\infty} H_i(N) S_N(i, j) H_j^\tau(N) \right\| \leq \lim_{N \rightarrow \infty} \frac{k_1^2}{N^\gamma} \sum_{i,j=0}^{\infty} e^{-k_2(i+j)} \|S_N(i, j)\| \xrightarrow{N \rightarrow \infty} 0.$$

Finally, we obtain the desired result

$$\begin{aligned} & \lim_{N \rightarrow \infty} \frac{\log^{1/4} N}{N} \sum_{n=1}^N (H_N(z) v_n)(H_N(z) v_n)^\tau \\ & = \lim_{N \rightarrow \infty} \frac{\log^{1/4} N}{N} \sum_{i,j=0}^{\infty} (H_i(N) S_N(i, j) H_j^\tau(N) \\ & \quad + \lim_{N \rightarrow \infty} \frac{\log^{1/4} N}{N} \sum_{i,j=0}^{\infty} H_i(N) \sum_{n=\max(i,j,1)}^N \delta_{ij} R_{n-i} H_j^\tau(N) \\ & = \lim_{N \rightarrow \infty} \frac{\log^{1/4} N}{N} \sum_{i=0}^N H_i(N) \sum_{n=\max(i,1)}^N R_{n-i} H_i^\tau(N) \\ & = \lim_{N \rightarrow \infty} \frac{\log^{1/4} N}{N} \left[ H_0(N) \sum_{n=1}^N R_n H_0^\tau(N) + \sum_{i=1}^N H_i(N) \sum_{n=0}^{N-i} R_n H_i^\tau(N) \right] \\ & = \lim_{N \rightarrow \infty} \frac{\log^{1/4} N}{N} \left[ H_0(N) \sum_{n=1}^N R_n H_0^\tau(N) \right. \\ & \quad \left. + \sum_{i=1}^N H_i(N) \left( \sum_{n=1}^N R_n + R_0 - \sum_{n=N-i+1}^N R_n \right) H_i^\tau(N) \right] \\ & = \lim_{N \rightarrow \infty} \frac{\log^{1/4} N}{N} \sum_{i=0}^N H_i(N) \sum_{n=1}^N R_n H_i^\tau(N). \quad \square \end{aligned}$$

*Proof of Lemma 6.* Set

$$S_N(i, j) \triangleq \sum_{n=1}^N w_{n+1-l-i} v_{n-m-j}^\tau.$$

Similar to the proof of (A.12), one can easily be convinced that

$$\lim_{N \rightarrow \infty} \frac{1}{N^\gamma} \sum_{i,j=0}^{\infty} H_i S_N(i, j) G_j^\tau = 0,$$

which is tantamount to (5.1).

Clearly, (5.2) can be verified in similar fashion.

By setting  $H_N(z) \equiv H(z)$  and  $v_n \equiv w_{n+1-l}$  in (A.12) we have

$$\lim_{N \rightarrow \infty} \frac{1}{N^\gamma} \sum_{i,j=0}^{\infty} H_i S_N(i,j) H_j^\tau = 0$$

where

$$S_N(i,j) = \sum_{n=\max(i,j,1)}^N [w_{n-l+1-i} w_{n-l+1-j}^\tau - \delta_{ij} R_{n-l+1-i}]$$

and

$$R_{n-l+1-i} \triangleq E w_{n-l+1-i} w_{n-l+1-i}^\tau.$$

Hence by the uniform boundedness of  $R_n$  we have

$$\begin{aligned} & \overline{\lim}_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N \|H(z) w_{n+1-l}\|^2 \\ &= \overline{\lim}_{N \rightarrow \infty} \text{tr} \left[ \frac{1}{N} \sum_{n=1}^N (H(z) w_{n+1-l})(H(z) w_{n+1-l})^\tau \right] \\ &= \overline{\lim}_{N \rightarrow \infty} \text{tr} \left[ \frac{1}{N} \sum_{i,j=0}^{\infty} H_i S_N(i,j) H_j^\tau + \sum_{i,j=0}^{\infty} H_i \frac{1}{N} \sum_{n=\max(i,j,1)}^N \delta_{ij} R_{n-l+1-i} H_j^\tau \right] \\ &= \overline{\lim}_{N \rightarrow \infty} \text{tr} \sum_{i=0}^{\infty} H_i \frac{1}{N} \sum_{n=\max(i,1)}^N R_{n-l+1-i} H_i^\tau < \infty. \end{aligned}$$

This completes the proof of the lemma.  $\square$

#### REFERENCES

- [1] K. J. ÅSTRÖM AND B. WITTENMARK, *On self-tuning regulators*, Automatica—J. IFAC, 9 (1973), pp. 185-195.
- [2] A. BECKER, P. R. KUMAR AND C. Z. WEI, *Adaptive control with the stochastic approximation algorithm—Geometry and convergence*, IEEE Trans. Automat. Control, AC-30 (1985), pp. 330-338.
- [3] P. E. CAINES AND S. LAFORTUNE, *Adaptive control with recursive identification for stochastic linear systems*, IEEE Trans. Automat. Control, AC-29 (1984), pp. 312-321.
- [4] P. E. CAINES AND S. MEYN, *On the zero divisor problem and singularities occurring in the recursive schemes of stochastic adaptive control*, Systems Control Lett., 6 (1985), pp. 235-238.
- [5] H. F. CHEN, *Quasi-least-squares identification and its strong consistency*, Internat. J. Control, 34 (1981), pp. 921-936.
- [6] ———, *Strong consistency of recursive identification under correlated noise*, J. Systems Sci. Math. Sci., 1 (1981), pp. 34-52.
- [7] ———, *Recursive system identification and adaptive control by use of the modified least squares algorithm*, this Journal, 22 (1984), pp. 758-776.
- [8] ———, *Recursive Estimation and Control for Stochastic Systems*, John Wiley, New York, 1985.
- [9] H. F. CHEN AND P. E. CAINES, *Strong consistency of the stochastic gradient algorithm of adaptive control*, IEEE Trans. Automat. Control, AC-30 (1985), pp. 189-192.
- [10] H. F. CHEN AND L. GUO (1984), *Adaptive control with recursive identification for stochastic linear systems*, in Advances in Control and Dynamic Systems, C. T. Leondes, ed., Vol. 24, Academic Press, New York, to appear.
- [11] ———, *Strong consistency of recursive identification by no use of persistent excitation condition*, Acta Math. Appl. Sinica, to appear.
- [12] ———, *Strong consistency of parameter estimates for discrete-time stochastic systems*, J. Systems Sci. Math. Sci., 5 (1985), pp. 81-93.

- [13] H. F. CHEN AND L. GUO, *The limit of stochastic gradient algorithm for identifying systems excited not persistently*, Kexue Tongbao (Science Bulletin), to appear.
- [14] Y. S. CHOW AND H. TEICHER (1978), *Probability Theory*, Springer, New York.
- [15] G. C. GOODWIN, P. T. RAMADGE AND P. E. CAINES, *Discrete-time stochastic adaptive control*, this Journal, 19 (1981), pp. 829–853.
- [16] T. KAILATH, *Linear Systems*, Prentice-Hall, Englewood Cliffs, NJ, 1980.
- [17] H. J. KUSHNER AND D. S. CLARK (1978), *Stochastic Approximation Methods for Constrained and Unconstrained Systems*, Springer, New York.
- [18] L. LJUNG, *Consistency of the least squares identification method*, IEEE Trans. Automat. Control, AC-22 (1976), pp. 551–575.
- [19] ———, *Analysis of recursive stochastic algorithms*, IEEE Trans. Automat. Control, AC-22 (1977), pp. 551–575.
- [20] J. B. MOORE, *On strong consistency of least squares identification algorithm*, Automatica—J. IFAC, 14 (1978), pp. 505–509.
- [21] K. S. SIN AND G. C. GOODWIN, *Stochastic adaptive control using a modified least squares algorithm*, Automatica—J. IFAC, 18 (1982), pp. 315–321.
- [22] V. SOLO, *The convergence of AML*, IEEE Trans. Automat. Control, AC-24 (1979), pp. 958–962.